

Written very late at night. Not yet checked. Very tired. Brain imploding. Do not invest money based on the calculations in this handout.

GEOMETRIC BROWNIAN MOTION AND THE BLACK-SCHOLES MODEL

From the handout on the Itô formula, you know that

GBM <1> 
$$S_t = \exp(\sigma B_t + \alpha t) \quad \text{with } B \text{ a Brownian motion}$$

then

$$S_t = 1 + \sigma S \bullet B_t + (a + \frac{1}{2}\sigma^2)S \bullet \mathcal{U}_t \quad \text{where } \mathcal{U}_t \equiv t.$$

In particular if we put  $\alpha = \mu - \frac{1}{2}\sigma^2$ , for a constant  $\mu$ , then

$$S_t = 1 + \sigma S \bullet B_t + \mu S \bullet \mathcal{U}_t,$$

or, in more traditional notation,

$$dS_t = \sigma S_t dB_t + \mu S_t dt \quad \text{or} \quad \frac{dS_t}{S_t} = \sigma dB_t + \mu dt.$$

Over a small time interval,  $[t, t + \delta]$  the proportional change  $\Delta S/S_t$  in  $S$  is approximately  $N(\mu\delta, \sigma^2\delta)$  distributed.

In the special case where  $\mu$  is zero,  $S_t = 1 + \sigma S \bullet B_t$ , which is a martingale.

The process  $S$  from <1> with  $\alpha = \mu - \frac{1}{2}\sigma^2$  is called a **geometric Brownian motion**. It is often used to model a stock price over time.

Notice that I have implicitly standardized the price so that  $S_0 = 1$ . In effect,  $S_t$  measures the price relative to the initial price. I will also make another standardization by assuming that the interest rate is zero, so that I don't have to discount future returns or introduce a bond into the calculations.

[§GHX] **1. Stochastic integrals with respect to stochastic integrals**

To apply the Itô formula in the arbitrage argument in the next section I will need a little piece of the calculus for stochastic integrals, namely

GHX <2> 
$$G \bullet (H \bullet X) = (GH) \bullet X.$$

Let me prove the equality only for elementary processes

$$G(t, \omega) = \sum_{i=0}^n g_i(\omega) \mathbb{I}\{t_i < t \leq t_{i+1}\}$$

$$H(t, \omega) = \sum_{i=0}^n h_i(\omega) \mathbb{I}\{t_i < t \leq t_{i+1}\}.$$

I postpone to a more rigorous course the formal passage to the limit from the elementary case to a case general enough to handle the process in the next section.

There is no loss of generality in assuming that both  $G$  and  $H$  are step functions for the same grid (we could always work with a common refinement of the grid for  $G$  and the grid for  $H$ ) and that we wish to establish equality <2> at a grid point (we could always add extra points to the grid). Writing  $Z$  for the process  $H \bullet X$ , we have  $Z_{t_k} = \sum_{i < k} h_i \Delta_i X$ , so that

$$\Delta_k Z = Z_{t_{k+1}} - Z_{t_k} = h_k \Delta_k X,$$

and

$$G \bullet Z_{t_k} = \sum_{i < k} g_i \Delta_i Z = \sum_{i < k} g_i h_i \Delta_i X = (GH) \bullet X_{t_k},$$

as asserted.



First you need to know what a change of measure is. If  $q(\omega)$  is a nonnegative random variable, living on a set  $\Omega$  already equipped with a probability, define

$$\mathbb{Q}A = \mathbb{E}_{\mathbb{P}}(q(\omega)\mathbb{I}\{\omega \in A\}) \quad \text{for } A \subseteq \Omega.$$

The subscript  $\mathbb{P}$  on the expectation will be needed to avoid confusion when we have more than one probability defined on  $\Omega$ . Provided  $\mathbb{E}_{\mathbb{P}}q = 1$ , the new  $\mathbb{Q}$  is a genuine probability. It satisfies all the usual properties, such as

$$\mathbb{Q}\left(\bigcup_{n \in \mathbb{N}} A_n\right) = \sum_{n \in \mathbb{N}} \mathbb{Q}(A_n) \quad \text{for disjoint events } \{A_n\}$$

and  $\mathbb{Q}\emptyset = 0$  and  $\mathbb{Q}\Omega = 1$ . More useful is the formula

$$\mathbb{E}_{\mathbb{Q}}X = \mathbb{E}_{\mathbb{P}}(Xq) \quad \text{for a random variable } X.$$

The random variable  $q$  is often called the density of  $\mathbb{Q}$  with respect to  $\mathbb{P}$ .

normalCoM    <6> **Example.** Suppose  $X_1, \dots, X_k$  are independent random variables with  $X_i \sim N(0, \sigma_i^2)$ , under the probability distribution  $\mathbb{P}$ . For arbitrary constants  $\{\alpha_i\}$  define a new probability,  $\mathbb{Q}$ , by means of the density

$$q = \exp\left(\sum_i (\alpha_i X_i - \frac{1}{2}\alpha_i^2 \sigma_i^2)\right).$$

Recall the formula  $\exp(\theta\mu + \frac{1}{2}\sigma^2\theta^2)$  for the moment generating function  $\mathbb{E}\exp(\theta Z)$  for a random variable  $Z$  with a  $N(\mu, \sigma^2)$  distribution. Together with the independence of the  $X_i$ 's under  $\mathbb{P}$ , this formula ensures that

$$\begin{aligned} \mathbb{E}_{\mathbb{P}}q &= \exp\left(-\sum_i \alpha_i^2 \sigma_i^2\right) \prod_i \mathbb{E}_{\mathbb{P}} \exp(\alpha_i X_i) \\ &= \exp\left(-\sum_i \alpha_i^2 \sigma_i^2\right) \prod_i \exp(\frac{1}{2}\alpha_i^2 \sigma_i^2) = 1. \end{aligned}$$

The  $\mathbb{Q}$  is a genuine probability distribution.

We can calculate the joint moment generating function of  $X_1, \dots, X_k$  under  $\mathbb{Q}$  by a similar argument. For constants  $\theta_1, \dots, \theta_k$ ,

$$\begin{aligned} \mathbb{E}_{\mathbb{Q}} \exp\left(\sum_i \theta_i X_i\right) &= \mathbb{E}_{\mathbb{P}} \exp\left(\sum_i \theta_i X_i + \sum_i (\alpha_i X_i - \frac{1}{2}\alpha_i^2 \sigma_i^2)\right) \\ &= \prod_i \mathbb{E}_{\mathbb{P}} \exp\left((\theta_i + \alpha_i)X_i - \frac{1}{2}\alpha_i^2 \sigma_i^2\right) \\ &= \prod_i \exp\left(\frac{1}{2}(\theta_i + \alpha_i)^2 \sigma_i^2 - \frac{1}{2}\alpha_i^2 \sigma_i^2\right) \\ &= \prod_i \exp\left(\frac{1}{2}\theta_i^2 \sigma_i^2 + \theta_i \alpha_i \sigma_i^2\right) \end{aligned}$$

The last expression is a product of moment generating functions for  $N(\alpha_i \sigma_i^2, \sigma_i^2)$  distributions. Under  $\mathbb{Q}$  the random variables  $X_1, \dots, X_k$  are still independent normals, with the same variances as under  $\mathbb{P}$ , but now the means have changed:  $\mathbb{E}_{\mathbb{Q}}X_i = \alpha_i \sigma_i^2$ . □

The change-of-measure trick also works for infinite collections of normally distributed random variables.

BMCoM    <7> **Example.** Suppose  $\{B_t : 0 \leq t \leq 1\}$  is a standard Brownian motion under  $\mathbb{P}$ . For a fixed constant  $\alpha$ , define

$$q(\omega) = \exp\left(\alpha B_1(\omega) - \frac{1}{2}\alpha^2\right)$$

Note that  $\mathbb{E}_{\mathbb{P}}q = 1$  because  $B_1 \sim N(0, 1)$  under  $\mathbb{P}$ .

If we change the measure to the  $\mathbb{Q}$  defined by the density  $q$  with respect to  $\mathbb{P}$ , we do not change the continuity of the sample paths of  $B$ . Suppose  $0 = t_0 < t_1 < \dots < t_{n+1} = 1$  is a grid with corresponding increments  $\Delta_i B$  for  $B$ . Under  $\mathbb{P}$  the increments are independent with  $\Delta_i B \sim N(0, \delta_i)$ , where  $\delta_i = t_{i+1} - t_i$ . The density  $q$  can also be written as

$$q = \exp\left(\sum_{i=0}^n (\alpha \Delta_i B - \frac{1}{2}\alpha^2 \delta_i)\right).$$

From Example <6> with  $\sigma_i^2 = \delta_i$ , deduce that the increments are again independent under  $\mathbb{Q}$  with  $\Delta_i B \sim N(\alpha\delta_i, \delta_i)$ , or  $\Delta_i B - \alpha\delta_i \sim N(0, \delta_i)$ . The process  $\tilde{B}_t = B_t - \alpha t$  has all the properties needed to characterize it as a Brownian motion under  $\mathbb{Q}$ . □

Now reconsider the stock prices modelled as a geometric Brownian motion,

$$S_t = \exp(\sigma B_t + (\mu - \frac{1}{2}\sigma^2)t) \quad \text{with } B \text{ a Brownian motion under } \mathbb{P}.$$

If  $\mathbb{Q}$  has density

$$q = \exp(-\mu B_1 - \frac{1}{2}\mu^2)$$

with respect to  $\mathbb{P}$  then  $\tilde{B}_t = B_t + \mu t$  is Brownian motion under  $\mathbb{Q}$  and

$$S_t = \exp(\sigma \tilde{B}_t - \frac{1}{2}\sigma^2 t) \quad \text{with } \tilde{B} \text{ a Brownian motion under } \mathbb{Q}.$$

With the change of measure we have effectively eliminated the drift coefficient  $\mu$ . Under  $\mathbb{Q}$ , the stock price is a martingale driven by the Brownian motion  $\tilde{B}$ .

Once again consider an option that promises to deliver a random amount  $Z$  at time  $t = 1$ . The variable could depend of the stock price history in a complicated way. For example, we could contemplate a most exotic option that delivers

$$Z = \max_{0 \leq t \leq 1} S_t - \sum_{j=107}^{233} S_{100/j}^2 \sin(S_{j/1000}) + \int_0^1 \cos(S_t^3) dt$$

at time  $t = 1$ . What matters most is that  $Z$  can also be thought of as a (weird) function of the  $\tilde{B}$  sample path: just insert  $\exp(\sigma \tilde{B}_t - \frac{1}{2}\sigma^2 t)$  wherever you see an  $S_t$  in the definition of  $Z$ , for various  $t$ . The function also depends on  $\sigma$ , but there is no  $\mu$  in sight.

The dramatic moment arrives.

Appeal to Fact <5> for the Brownian motion  $\tilde{B}$  to express  $Z$  as

$$Z = C + H \bullet \tilde{B}_1$$

for some constant  $C$  and some adapted process  $H$ . (Maybe you should check that  $\mathbb{E}_{\mathbb{Q}} Z^2 < \infty$  for the  $Z$  you have in mind.) If we could trade directly in  $\tilde{B}$ , we could interpret  $H \bullet \tilde{B}$  as a trading scheme. We need to convert to a scheme trading in the stock price by means of the representation

$$S_t = 1 + \sigma S \bullet \tilde{B}_t \quad \text{under } \mathbb{Q}.$$

The equality <2> again comes to the rescue, if we integrate the process  $1/S_t$  with respect to the processes on both sides of the previous display.

$$\begin{aligned} (1/S) \bullet S_t &= (1/S) \bullet 1_t + (1/S) \bullet (\sigma S \bullet \tilde{B}_t) \\ &= 0 + \sigma (S/S) \bullet \tilde{B}_t \quad \text{cf. increments of a constant process} \\ &= \sigma \tilde{B}_t. \end{aligned}$$

Similarly,

$$H \bullet \tilde{B}_t = \frac{1}{\sigma} H \bullet ((1/S) \bullet S)_t = \frac{1}{\sigma} (H/S) \bullet S_t.$$

Write  $K_t$  for  $(1/\sigma)(H/S)_t$ . Then we have a trading scheme to recover the amount  $Z - C$  at time  $t = 1$ :

$$Z = C + K \bullet S_1$$

You might be a bit disappointed that you know only how to trade under  $\mathbb{Q}$  if in fact you live in the world where  $\mathbb{P}$  is in control and  $S$  is not a martingale because of that pesky, unknown  $\mu$ . (You did say that you knew the value of  $\sigma$ , didn't you?)

Not to worry. Think of  $K$  as a shorthand for a sequence of elementary processes,

$$K_n(t) = \sum_j k_{n,j}(\omega) \mathbb{I}\{t_{n,j} < t \leq t_{n,j+1}\}$$

for which  $K_n \bullet S$  converges to  $K \bullet S$ . The trading scheme  $K_n$  can be spelled out as

for each  $j$ : buy  $k_{n,j}$  shares at time  $t_{n,j}$  then sell them at time  $t_{n,j+1}$

At this point I need to be a little more precise about the sense of the convergence. In fact, I need (and the stochastic calculus gives) convergence in  $\mathbb{Q}$  probability:

$$\mathbb{Q}\{|K_n \bullet S - K \bullet S| > \epsilon\} \rightarrow 0 \quad \text{as } n \rightarrow \infty, \text{ for each } \epsilon > 0.$$

The nice thing about the relationship between  $\mathbb{Q}$  and  $\mathbb{P}$  is: sequences that converge in  $\mathbb{Q}$ -probability also converge (to the same thing) in  $\mathbb{P}$ -probability. The idealized trading scheme  $K$  is a limit of the elementary schemes  $K_n$  under both  $\mathbb{Q}$  and  $\mathbb{P}$ .

Some calculations needed here.

Before we leave the  $\mathbb{Q}$ -world, note that  $S$  and  $K \bullet S$  are both martingales under  $\mathbb{Q}$ . In particular,

$$0 = \mathbb{E}_{\mathbb{Q}} K \bullet S_0 = \mathbb{E}_{\mathbb{Q}} K \bullet S_1$$

and hence

$$\mathbb{E}_{\mathbb{Q}} Z = C,$$

a calculation that we could, in principle, carry out.

Back in the world controlled by  $\mathbb{P}$ , we therefore have a trading scheme,  $K$ , that delivers the amount  $Z - C$  at time  $t = 1$ . We should pay  $C$  at time  $t = 0$  to receive  $Z$  at time  $t = 1$ .

In short: to find the price to pay at time  $t = 0$  for receiving  $Z$  at time  $t = 1$ ,

- (i) Find the probability measure  $\mathbb{Q}$  that makes  $S$  a  $\mathbb{Q}$ -martingale.
- (ii) Hope (or invoke some probability theorem to show) that convergence in  $\mathbb{Q}$ -probability is the same as convergence in  $\mathbb{P}$ -probability.
- (iii) Calculate the price as  $C = \mathbb{E}_{\mathbb{Q}} Z$ .

**BScall** <8> **Example.** Suppose  $Z = (S_1 - K)^+$ , which I believe is the return from the option known as a call with strike price  $K$ . Calculate.

$$C = \mathbb{E}_{\mathbb{Q}} \exp(\sigma S_1 - K)^+ = \mathbb{E}_{\mathbb{Q}} \left( \exp(\sigma \tilde{B}_1 - \frac{1}{2}\sigma^2) - K \right)^+.$$

Under  $\mathbb{Q}$ , the random variable  $W = \tilde{B}_1$  has a standard normal distribution. Also

$$\exp(\sigma W - \frac{1}{2}\sigma^2) \geq K \text{ if and only if } W \geq L := \frac{1}{\sigma} \log K + \frac{1}{2}\sigma$$

Write  $\bar{\Phi}(t) = 1 - \Phi(t)$  for the standard normal tail probability. Calculate.

$$\begin{aligned} C &= \mathbb{E}_{\mathbb{Q}} \left( \exp(\sigma W - \frac{1}{2}\sigma^2) - K \right) \mathbb{I}\{W \geq L\} \\ &= \frac{1}{\sqrt{2\pi}} \int_L^\infty \exp(\sigma x - \frac{1}{2}\sigma^2 - \frac{1}{2}x^2) dx - K \mathbb{E}_{\mathbb{Q}} \mathbb{I}\{W \geq L\} \\ &= \frac{1}{\sqrt{2\pi}} \int_L^\infty \exp(-\frac{1}{2}(x - \sigma)^2) dx - K \bar{\Phi}(L) \\ &= \bar{\Phi}(L - \sigma) - K \bar{\Phi}(L) \end{aligned}$$

I sure hope the last expression agrees with the textbooks for the case where  $S_0 = 1$  and there is a zero interest rate. Stay tuned for the corrected version with the correct result. □

#### REFERENCES

- Pollard, D. (2001), *A User's Guide to Measure Theoretic Probability*, Cambridge University Press.
- Wilmott, P., Howison, S. & Dewynne, J. (1995), *The Mathematics of Financial Derivatives: a Student Introduction*, Cambridge University Press.