

Maximum likelihood in an infinite-dimensional exponential family

David Pollard

Yale University

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Based on:

Dou, Pollard, and Zhou (2010) "Functional regression for general exponential families"
[arXiv:1001.3742v1 \[math.ST\]](https://arxiv.org/abs/1001.3742v1) 21 Jan 2010

Slides for this talk at: www.stat.yale.edu/~pollard/Talks

1. Problem: estimate unknown β

- ▶ Observe independent $(y_1, x_1), (y_2, x_2), \dots$ with $y_i | x_i \sim Q_{\theta_i}$ and (nonrandom or condition) $x'_i = (x_{i1}, x_{i2}, \dots)$ and $\theta_i = \langle x_i, \beta \rangle$



$$\frac{dQ_{\theta}}{dQ_0} = \exp(y\theta - \Psi(\theta)) \quad \theta \in \mathbb{R}$$

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Example: for $Q_{\theta} = \text{Bernoulli} \left(\frac{e^{\theta}}{1 + e^{\theta}} \right)$ on $\{0, 1\}$

$$\Psi(\theta) = \log \left(\frac{1 + e^{\theta}}{2} \right)$$

Example: for $Q_{\theta} = \text{Poisson} (e^{\theta})$ on $\{0, 1, 2, \dots\}$

$$\Psi(\theta) = e^{\theta} - 1$$

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$$\frac{dQ_{\theta}}{dQ_0} = \exp(y\theta - \Psi(\theta)) \quad \theta \in \mathbb{R}$$

- ▶ Ψ convex with $\Psi(0) = 0$
- ▶ MLE: maximize a concave function of b ?

$$\hat{b} \stackrel{?}{=} \operatorname{argmax}_{b \in ?} \sum_{i \leq n} (y_i \langle x_i, b \rangle - \Psi(\langle x_i, b \rangle))$$

Component of \hat{b} in $\operatorname{span}\{x_1, \dots, x_n\}^{\perp}$ is unconstrained.

2. Simplify: $x_i \in \mathbb{R}^N$, with $N = N_n$?

► MLE

$$\hat{b}_n = \operatorname{argmax}_{b \in \mathbb{R}^N} \sum_{i \leq n} (y_i x_i' b - \Psi(x_i' b)) \quad (\text{concave})$$

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► $\hat{b}_n = \beta + J_n^{-1/2} \hat{t}_n$ where $\hat{t}_n = \underset{t \in \mathbb{R}^N}{\operatorname{argmax}} L_n(t)$ for

$$L_n(t) = \sum_{i \leq n} (y_i w_i' t - \Psi(\theta_i + w_i' t)) \quad (\text{concave})$$

3. Why reparametrize?

- ▶ [from last slide] $\hat{t}_n = \operatorname{argmax}_{t \in \mathbb{R}^N} L_n(t)$ where

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$$\dot{L}(t) = \frac{\partial}{\partial t} L_n(t)$$

$$= \sum_{i \leq n} \left(y_i - \dot{\Psi}(\theta_i) \right) w_i' - \sum_{i \leq n} \left(\dot{\Psi}(\theta_i + w_i' t) - \dot{\Psi}(\theta_i) \right) w_i'$$

$$= T_n' - t' \sum_{i \leq n} w_i w_i' \ddot{\Psi}(\theta_i) - R_n(t)'$$

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- ▶ Should have $\hat{t}_n \approx T_n$ (even when $N_n \rightarrow \infty$)

4. Remainder term in $\dot{L}_n(t) = T'_n - t' - R_n(t)'$

► Assume

$$|\overset{\bullet\bullet\bullet}{\Psi}(\theta + h)| \leq \overset{\bullet\bullet}{\Psi}(\theta)G(|h|) \quad \text{for all } \theta \text{ and } h$$

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- ▶ With $M_n := \max_{i \leq n} |w_i|$,

$$\begin{aligned} |R_n(t)| &= \left| \sum_{i \leq n} \left(\overset{\bullet}{\Psi}(\theta_i + w'_i t) - \overset{\bullet}{\Psi}(\theta_i) - w'_i t \overset{\bullet\bullet}{\Psi}(\theta_i) \right) w'_i \right| \\ &\leq \frac{1}{2} M_n G(M_n |t|) t' \sum_{i \leq n} w_i w'_i \overset{\bullet\bullet\bullet}{\Psi}(\theta_i) t \end{aligned}$$

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5. Directional derivatives for $\dot{L}_n(t) = T'_n - t' - R_n(t)'$

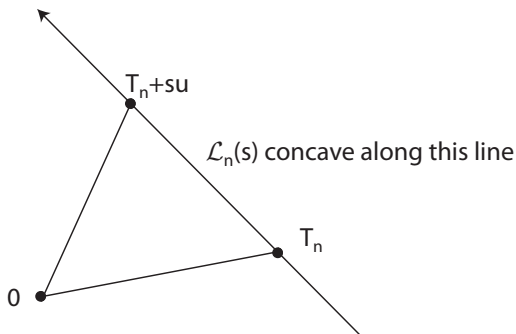
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$$|R_n(T_n \pm \epsilon u)'u| \leq \frac{1}{2} M_n G(M_n |T_n \pm \epsilon u|) |T_n \pm \epsilon u|^2$$

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if

$$M_n \leq \frac{\epsilon^2}{2G(1)N} \quad \text{and} \quad |T_n|^2 \leq N/\epsilon$$

6. Maximization of concave $\mathcal{L}_n(s) = L_n(T_n + su)$

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and

$$|J_n^{1/2}(\hat{b}_n - \beta) - T_n| = |\hat{t}_n - T_n| \leq \epsilon$$

7. General case: $\theta_i = \langle x_i, \beta \rangle$ with $x_i \in \mathbb{R}^N$

- ▶ Truncate $z = (z_1, z_2, \dots) = \pi_N z + \pi_N^\perp z$:

$$\pi_N z = (z_1, \dots, z_N, 0, 0, \dots), \quad \pi_N^\perp z = (0, \dots, 0, z_{N+1}, z_{N+2}, \dots)$$

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- ▶ (Conditional) distributions for y_1, \dots, y_n :

$$\mathbb{P}_n = \otimes_{i \leq n} Q_{\theta_i}, \quad Q_{n,N} = \otimes_{i \leq n} Q_{\theta_i^*}$$

Bias term δ_i for $\hat{b}_{n,N}$ under \mathbb{P}_n ; no bias under $Q_{n,N}$

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- ▶ Theory from N_n case (as before) gives

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$$\begin{aligned} H^2(\mathbb{P}_n, Q_{n,N}) &= \sum_{i \leq n} H^2(Q_{\theta_i}, Q_{\theta_i^*}) \\ &\leq \sum_{i \leq n} \delta_i^2 \ddot{\Psi}(\theta_i) (1 + |\delta_i|) G(|\delta_i|) \end{aligned}$$

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- ▶ Assumptions on N_n and on decay of β_j 's needed to make $\|\mathbb{P}_n - Q_{n,N}\|_{TV} \rightarrow 0$ and $|\pi_N^\perp \beta|$ suitably small.
- ▶ Deduce $\mathbb{P}_n\{\widehat{b}_{n,N} \approx \pi_N \beta\} \approx 1$ if

References

-  Hall, P. and J. L. Horowitz (2007).
Methodology and convergence rates for functional linear regression.
Annals of Statistics 35(1), 70–91.
-  Hjort, N. L. and D. Pollard (1993).
Asymptotics for minimisers of convex processes.
Technical report, Yale University.