#### S&DS 241 Lecture 7 Union bound. Inclusion-Exclusion principles.

B-H: 1.6,4.4

Recall: Union of two events

From axioms of probability:

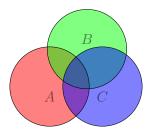
$$P(A \cup B) = P(A) + P(B) - P(AB)$$

(Again: we omit  $\cap$  and write AB for intersection)



Recall: Union of three events

$$P(A \cup B \cup C) = P(A) + P(B) + P(C)$$
$$- P(AB) - P(BC) - P(CA)$$
$$+ P(ABC)$$





How to deal with

 $P\left(A_1 \cup A_2 \cup \cdots \cup A_n\right)$ 

### Today

How to deal with

$$P\left(A_1 \cup A_2 \cup \cdots \cup A_n\right)$$

We will learn:

- 1 Inequality: union bound (Boole's or Bonferroni's inequality)
- 2 Equality: inclusion-exclusion principle

Two events:

#### $P\left(A \cup B\right) \le P\left(A\right) + P\left(B\right)$

**Proof**:  $P(A \cup B) = P(A) + P(B) - P(AB)$ 

• Two events:

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**Proof**:  $P(A \cup B) = P(A) + P(B) - P(AB)$ 

• Corollary:

 $P(A_1 \cup \ldots \cup A_n) \le P(A_1) + \ldots + P(A_n)$ 

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#### Remarks

• Convenient to use: no need to deal with intersections or postulate independence

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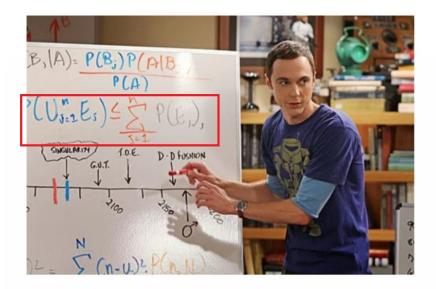
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• Corollary:

 $P(A_1 \cup \ldots \cup A_n) \le P(A_1) + \ldots + P(A_n)$ 

#### Remarks

- Convenient to use: no need to deal with intersections or postulate independence
- Conservative estimate and can be useless (might exceed 1)



A student takes 4 classes; each fails with probability 3%. Consider

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 $P(A_1 \cup A_2 \cup A_3 \cup A_4) = 1 - P(A_1^c A_2^c A_3^c A_4^c) = 1 - (1 - 3\%)^4 \approx 11.5\%$ 

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 Independence might not be realistic to assume. But union bound always applies.

$$P(A_1 \cup A_2 \cup \dots \cup A_n)$$
  
=  $\sum_{i=1}^n P(A_i)$   
-  $\sum_{1 \le i < j \le n}^n P(A_i A_j)$   
+  $\sum_{1 \le i < j < k \le n}^n P(A_i A_j A_k)$   
-  $\dots$   
+  $(-1)^{n-1} P(A_1 A_2 \cdots A_n).$ 

$$P(A_{1} \cup A_{2} \cup \dots \cup A_{n})$$

$$= \sum_{i=1}^{n} P(A_{i}) \qquad n \text{ terms}$$

$$- \sum_{1 \le i < j \le n}^{n} P(A_{i}A_{j}) \qquad \text{all } \binom{n}{2} \text{ pairs}$$

$$+ \sum_{1 \le i < j < k \le n}^{n} P(A_{i}A_{j}A_{k}) \qquad \text{all } \binom{n}{3} \text{ triples}$$

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**Proof** 1 induction on n (exercise)

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Proof 1 induction on n (exercise) Proof 2 indicator random variables (later)

A hat-checker in a restaurant, having checked n hats, gets them hopelessly scrambled and returns them at random to the guests as they leave. Let X = number of guests with own hats. Find E(X) without finding PMF.

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• Note  $X = X_1 + \cdots + X_n$  and hence

$$E(X) = E(X_1) + \dots + E(X_n) = P(A_1) + \dots + P(A_n) = \frac{1}{n} + \dots + \frac{1}{n} = 1_{\frac{9}{24}}$$

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Recall

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Note

 $\{X = 0\} = \{\text{nobody gets own hat}\} = A_1^c \cap A_2^c \dots \cap A_n^c$  $\{X > 0\} = \{\text{somebody gets own hat}\} = A_1 \cup A_2 \dots \cup A_n$ 

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• Plan: use inclusion-exclusion principle to find  $P(A_1 \cup A_2 \cdots \cup A_n)$ .

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- Plan: use inclusion-exclusion principle to find  $P(A_1 \cup A_2 \cdots \cup A_n)$ .
- Union bound is not useful here:

$$P(A_1 \cup A_2 \cdots \cup A_n) \le P(A_1) + P(A_2) + \cdots + P(A_n) = 1$$

$$P(A_{1} \cup A_{2} \cup \dots \cup A_{n})$$

$$= \sum_{i=1}^{n} P(A_{i}) \qquad n \text{ terms}$$

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$$\begin{split} &P\left(A_1 \cup A_2 \cup \dots \cup A_n\right) \\ &= \sum_{i=1}^n P\left(A_i\right) & n \text{ terms} \\ &- \sum_{1 \leq i < j \leq n}^n P\left(A_i A_j\right) & \text{all } \binom{n}{2} \text{ pairs} \\ &+ \sum_{1 \leq i < j < k \leq n}^n P\left(A_i A_j A_k\right) & \text{all } \binom{n}{3} \text{ triples} \\ &- \dots \\ &+ \left(-1\right)^{n-1} P\left(A_1 A_2 \dots A_n\right). \end{split}$$

We already know:  $P(A_i) = 1/n$ ,

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We already know:  $P(A_i) = 1/n$ , or another way:  $P(A_i) = \frac{(n-1)!}{n!}$ 

### Pairwise intersection

$$P(A_i A_j) = P(A_i)P(A_j | A_i) = \frac{1}{n} \times \frac{1}{n-1}$$

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Total number of pairs:  $\binom{n}{2} = \frac{n!}{2!(n-2)!}$ 

# Triple intersection

$$P(A_i A_j A_k) = P(A_i) P(A_j | A_i) P(A_k | A_i A_j) = \frac{1}{n} \times \frac{1}{n-1} \times \frac{1}{n-2}$$

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Total number of pairs:  $\binom{n}{3} = \frac{n!}{3!(n-3)!}$ 

# Intersection of all

$$P\left(A_1 A_2 \cdots A_n\right) = \frac{1}{n!}$$

# Inclusion-exclusion principle

$$\begin{split} &P\left(A_{1}\cup A_{2}\cup\cdots\cup A_{n}\right)\\ &=\sum_{i=1}^{n}P\left(A_{i}\right) & n \text{ terms}\\ &-\sum_{1\leq i< j\leq n}^{n}P\left(A_{i}A_{j}\right) & \text{all } \binom{n}{2} \text{ pairs}\\ &+\sum_{1\leq i< j< k\leq n}^{n}P\left(A_{i}A_{j}A_{k}\right) & \text{all } \binom{n}{3} \text{ triples}\\ &-\cdots\\ &+\left(-1\right)^{n-1}P\left(A_{1}A_{2}\cdots A_{n}\right). \end{split}$$

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$$-\dots$$

$$+(-1)^{n-1}\frac{1}{n!}$$

#### Inclusion-exclusion principle

$$P(A_{1} \cup A_{2} \cup \dots \cup A_{n})$$
=1
$$-\binom{n}{2} \times \frac{(n-2)!}{n!} + \binom{n}{3} \times \frac{(n-3)!}{n!} + (-1)^{n-1} \frac{1}{n!} = 1 - \frac{1}{2!} + \frac{1}{3!} - \dots + (-1)^{n-1} \frac{1}{n!}$$

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$$P(X = 0)$$
  
=1 - P (A<sub>1</sub> \cup A<sub>2</sub> \cup \dots \cup A<sub>n</sub>)  
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If there are many people,  $P({\rm nobody\ gets\ own\ hat})\approx 1/e$  Why?

• Taylor expansion: [B-H: Math Appendix A.8.3]

$$e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \cdots$$

• Expand at x = -1:  $e^{-1} = 1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \cdots$ 

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• Expand at x = -1:  $e^{-1} = 1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \cdots$ What about P(X = 1)?

# Is 1/e surprising?

• Let's suppose, hypothetically,  $A_i$ 's were independent. Then the probability of nobody gets own hat would be

$$\left(1-\frac{1}{n}\right)^n \xrightarrow{n \to \infty} \frac{1}{e}$$

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• 
$$P(A_2|A_1) = \frac{1}{n-1} > P(A_2) = \frac{1}{n}$$

 Nevertheless, the dependence is rather weak when n is large, and inclusion-exclusion principle makes it possible to rigorously compute it.

#### Let's prove the inclusion-exclusion principle

### Recall: indicator random variable

• For every event A, we define a binary-valued random variable:

$$\mathbf{1}_{A} = \begin{cases} 1 & A \text{ occurs} \\ 0 & A \text{ does not occur} \end{cases}$$

• Important relation ("fundamental bridge"):

 $E(\mathbf{1}_A) = P(A)$ 

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• Calculus of indicators (shorthand:  $AB = A \cap B$ )

$$\mathbf{I}_A \times \mathbf{1}_B = \mathbf{1}_{AB}$$

▶  $1_{A^c} = 1 - 1_A$ 

 $\mathbf{1}_{A_1\cup A_2\cup\cdots\cup A_n}$ 

$$\mathbf{1}_{A_1\cup A_2\cup\cdots\cup A_n}=1-\mathbf{1}_{A_1^cA_2^c\cdots A_n^c}$$

$$\mathbf{1}_{A_1 \cup A_2 \cup \dots \cup A_n} = 1 - \mathbf{1}_{A_1^c A_2^c \cdots A_n^c}$$
$$= 1 - \mathbf{1}_{A_1^c} \times \mathbf{1}_{A_2^c} \times \dots \times \mathbf{1}_{A_n^c}$$

$$\begin{aligned} \mathbf{1}_{A_{1}\cup A_{2}\cup\cdots\cup A_{n}} =& 1 - \mathbf{1}_{A_{1}^{c}A_{2}^{c}\cdots A_{n}^{c}} \\ =& 1 - \mathbf{1}_{A_{1}^{c}}\times\mathbf{1}_{A_{2}^{c}}\times\cdots\times\mathbf{1}_{A_{n}^{c}} \\ =& 1 - (1 - \mathbf{1}_{A_{1}})(1 - \mathbf{1}_{A_{2}})\cdots\cdots(1 - \mathbf{1}_{A_{n}}) \end{aligned}$$

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Take expectation on both sides and invoke linearity of expectation:

$$E(\mathbf{1}_{A_1\cup A_2\cup\cdots\cup A_n}) = \sum_{i=1}^n E(\mathbf{1}_{A_i})$$
  
$$-\sum_{1\leq i< j\leq n}^n E(\mathbf{1}_{A_iA_j})$$
  
$$+\sum_{1\leq i< j< k\leq n}^n E(\mathbf{1}_{A_iA_jA_k})$$
  
$$-\cdots$$
  
$$+(-1)^{n-1} E(\mathbf{1}_{A_1A_2\cdots A_n})$$

Thus

$$P(A_1 \cup A_2 \cup \dots \cup A_n) = \sum_{i=1}^n P(A_i)$$
  
$$-\sum_{1 \le i < j \le n}^n P(A_i A_j)$$
  
$$+\sum_{1 \le i < j < k \le n}^n P(A_i A_j A_k)$$
  
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- Taking expectations on both sides and applying linearity  $\implies$

$$P(A_1 \cup \ldots \cup A_n) \le P(A_1) + \ldots + P(A_n)$$

• Using similar reasoning, we can show (Exercise)

$$P(A_1 \cup \ldots \cup A_n) \ge P(A_1) + \ldots + P(A_n) - \sum_{i < j} P(A_i A_j)$$