S&DS 241 Lecture 9 Poisson distribution

B-H: 4.7,4.8, math appendix A.8,A.9

Poisson distribution

A random variable X is said to have a Poisson distribution with parameter $\lambda \ge 0$, denoted by $X \sim \text{Pois}(\lambda)$, if¹

$$P(X = k) = \frac{e^{-\lambda}\lambda^k}{k!}, \quad k = 0, 1, 2, \dots$$

i.e.,
$$P(X=0)=e^{-\lambda},$$
 $P(X=1)=\lambda e^{-\lambda},$ etc.



¹As a convention 0! = 1.

Poisson distribution



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Siméon Denis Poisson

Application of Poisson distribution

Poisson distribution arises in studying many rare events

- Number of photons arriving at a detector in a fixed interval of time.
- Telephone traffic to base station / arrival of spam emails
- Modeling insurance claims
- Typos in a large document

Historical examples

 Abraham de Moivre (1711) and Simon Denis Poisson (1837): Theorized about the number of wrongful convictions in a given country during a time interval of given length



²For dataset see Table 5.5 of Grinstead-Snell.

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 Ladislaus von Bortkiewicz (1898): Investigated the number of soldiers in the Prussian army killed accidentally by horse kicks in 1875-1894.²

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Relations between Poisson and Binomial distributions

- Poisson distribution is a good approximation of Bin(n, p) when n is large and p is small.
- Specifically,
 - when n is large (many independent trials)
 - \blacktriangleright success probability p is inversely proportional to n

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Question

When n is large (partition by milliseconds, nanoseconds, etc), how does X behave?























Bin(100, 4/100) vs Pois(4)

• Precise statement:

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in the sense of convergence of PMF: for any fixed k,

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• Example (*k* = 0):

$$P\left(\mathsf{Bin}\left(n,\lambda/n\right)=0\right) = \left(1-\frac{\lambda}{n}\right)^n \to e^{-\lambda} = P(\mathsf{Pois}(\lambda)=0)$$

Fact from calculus [B-H A.2.5]: $\lim_{n\to\infty} (1+\frac{x}{n})^n = e^x$

In general: for fixed $k \ge 0$,

$$P\left(\operatorname{Bin}\left(n,\lambda/n\right)=k\right)$$

$$= \binom{n}{k} \left(\frac{\lambda}{n}\right)^{k} \left(1-\frac{\lambda}{n}\right)^{n-k}$$

$$= \frac{n\left(n-1\right)\cdots\cdots\left(n-k+1\right)}{k!} \left(\frac{\lambda}{n}\right)^{k} \left(1-\frac{\lambda}{n}\right)^{n-k}$$

$$= \frac{\lambda^{k}}{k!} \cdot 1 \cdot \underbrace{\left(1-\frac{1}{n}\right)}_{\rightarrow 1} \cdots \underbrace{\left(1-\frac{k-1}{n}\right)}_{\rightarrow 1} \underbrace{\left(1-\frac{\lambda}{n}\right)^{-k}}_{\rightarrow 1} \underbrace{\left(1-\frac{\lambda}{n}\right)^{n}}_{\rightarrow e^{-\lambda}}$$

$$\rightarrow \frac{e^{-\lambda}\lambda^{k}}{k!}$$

Properties of Poisson

Verify normalization of Poisson PMF

• Recall: Taylor expansion

$$e^{\lambda} = 1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \dots + \frac{\lambda^k}{k!} + \dots$$
$$= \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}$$

• Hence

$$\sum_{k=0}^{\infty} \frac{e^{-\lambda} \lambda^k}{k!} = 1$$

Expectation

• $X \sim \mathsf{Pois}(\lambda)$. Then

$$E(X) = \lambda,$$

Intuition: the expectation of of $Bin(n, \frac{\lambda}{n})$ is λ , for any n.

• Direct verification:

$$E(X) = \sum_{k=0}^{\infty} k \frac{e^{-\lambda} \lambda^k}{k!} = \sum_{k=1}^{\infty} k \frac{e^{-\lambda} \lambda^k}{k!}$$
$$= \sum_{k=1}^{\infty} \frac{e^{-\lambda} \lambda^k}{(k-1)!}$$
$$= \sum_{j=0}^{\infty} \frac{e^{-\lambda} \lambda^{j+1}}{j!}$$
$$= \lambda \sum_{j=0}^{\infty} \frac{e^{-\lambda} \lambda^j}{j!} = \lambda$$









Sum of independent Poissons

Let $X \sim \mathsf{Pois}(\lambda)$ and $Y \sim \mathsf{Pois}(\mu)$ be independent. Then

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• This property is also inherited from Binomials: when $p \rightarrow 0$,

$$\begin{split} & \mathsf{Bin}\Big(\frac{\lambda}{p},p\Big) \to \mathsf{Pois}(\lambda) \\ & \mathsf{Bin}\Big(\frac{\mu}{p},p\Big) \to \mathsf{Pois}(\mu) \end{split}$$

and from last lecture we know that

$$\underbrace{\mathsf{Bin}\Big(\frac{\lambda}{p},p\Big) + \mathsf{Bin}\Big(\frac{\mu}{p},p\Big)}_{\text{independent}} = \mathsf{Bin}\Big(\frac{\lambda+\mu}{p},p\Big) \to \mathsf{Pois}(\lambda+\mu)$$

Direct verification

Let $X \sim \mathsf{Pois}(\lambda)$ and $Y \sim \mathsf{Pois}(\mu)$ be independent.

$$\begin{split} P\left(X+Y=k\right) &= \sum_{j=0}^{k} P\left(X=j\right) P\left(Y=k-j\right) \\ &= \sum_{j=0}^{k} \frac{e^{-\lambda}\lambda^{j}}{j!} \frac{e^{-\mu}\mu^{k-j}}{(k-j)!} \\ &= \frac{e^{-(\lambda+\mu)}}{k!} \sum_{j=0}^{k} k! \frac{\lambda^{j}}{j!} \frac{\mu^{k-j}}{(k-j)!} \\ &= \frac{e^{-(\lambda+\mu)}}{k!} \sum_{j=0}^{k} \binom{k}{j} \lambda^{j} \mu^{k-j} \\ &= \frac{e^{-(\lambda+\mu)}}{k!} \left(\lambda+\mu\right)^{k} \end{split}$$
 binomial

expansion (Lec 8)

Conditioned on the sum

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$$P(X = j | X + Y = k)$$

$$= \frac{P(X = j, X + Y = k)}{P(X + Y = k)} = \frac{P(X = j)P(Y = k - j)}{P(X + Y = k)}$$

$$= \frac{\frac{e^{-\lambda_{\lambda}j}}{j!} \frac{e^{-\mu}\mu^{k-j}}{(k-j)!}}{\frac{e^{-(\lambda+\mu)}}{k!} (\lambda+\mu)^{k}} = \binom{k}{j} \left(\frac{\lambda}{\lambda+\mu}\right)^{j} \left(\frac{\mu}{\lambda+\mu}\right)^{k-j}$$

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Conditioned on the sum X + Y = k, X is distributed as $Bin(k, \frac{\lambda}{\lambda+\mu})$.

- Replace Bin(n, p) by Pois(np).
- We know this is accurate when $n \to \infty$ and np converges to a constant.

Poisson approximation

Poisson paradigm

Let A_1, \ldots, A_n be a collection of events (e.g. indexed by time or space)

- n is large
- $p_j = P(A_j)$ small
- A_j are (approximately) independent

Let $X = \sum_{j=1}^{n} \mathbf{1}_{A_j}$ count how many of the events A_j 's occur. Then X is approximately Poisson distributed as $\mathsf{Pois}(\sum_{j=1}^{n} p_j)$

Example

Consider the statistics of flying bomb hits in the south of London during World War II.

- The entire area is divided into a grid of N=576 small areas of size $\frac{1}{4}{\rm km}^2$ each.
- The total number of hits is 537.
- The average number of hits per square is 537/576 = 0.93 hits per square.



Grinstead-Snell Example 5.4 and Feller vol I Sec VI.7(b).

London bombing data fitted by Poisson

| Number of hits k | 0 | 1 | 2 | 3 | 4 | 5 + |
|---|-----|-----|----|----|---|-----|
| Number of areas with k hits | 229 | 211 | 93 | 35 | 7 | 1 |
| 576 · $P\left(X=k\right)$, $X\sim Pois\left(0.93\right)$ | 227 | 211 | 99 | 31 | 7 | 2 |



Figure: Empirical proportions vs. Pois(0.93)

Interpretation

VI.71 161 OBSERVATIONS FITTING THE POISSON DISTRIBUTION

Poisson distribution is surprisingly good; as judged by the χ^2 -criterion, under ideal conditions some 88 per cent of comparable observations should show a worse agreement. It is interesting to note that most people believed in a tendency of the points of impact to cluster. If this were true, there would be a higher frequency of areas with either many hits or no hit and a deficiency in the intermediate classes. Table 4 indicates perfect randomness and homogeneity of the area; we have here an instructive illustration of the established fact that to the untrained eye randomness appears as regularity or tendency to cluster.

| EXAMPLE (b): FLYING-BOMB HITS ON LONDON | | | | | | | | |
|---|--------|--------|-------|-------|------|------------|--|--|
| k | 0 | 1 | 2 | 3 | 4 | 5 and over | | |
| N_k | 229 | 211 | 93 | 35 | 7 | 1 | | |
| Np(k; 0.9323) | 226.74 | 211.39 | 98.54 | 30.62 | 7.14 | 1.57 | | |

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