S&DS 241 Lecture 14

Uniform distribution, exponential distribution, random number generation

B-H 5.2,5.3,5.5

Discrete uniform distribution

"Uniform" means "equally likely".

• Fair die: PMF is flat



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• How to extend this to continuous random variables?

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• How to extend this to continuous random variables? PDF is flat



Uniform distribution

A continuous random variable X is said to be uniformly distributed in the interval (a, b), denoted by $X \sim \text{Unif}(a, b)$, if it has the following PDF:

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a < x < b\\ 0 & \text{else} \end{cases}$$



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- Continuous analog of "equally likely"
- Support of X: [a, b]
- CDF:

$$F_X(x) = \begin{cases} 0 & x \le a \\ \frac{x-a}{b-a} & a < x < b \\ 1 & x \ge b \end{cases}$$



Example from last lecture

• Angle of a randomly spinned wheel



Then $X \sim \text{Unif}(0, 2\pi)$

 $\mathsf{Unif}(0,1)$

PDF $f_X(x)$ 1 $f_X(x) = \begin{cases} 1 & 0 < x < 1 \\ 0 & \text{else} \end{cases}$ xÔ 1 CDF: $F_X(x)$ $F_X(x) = \begin{cases} 0 & x \le 0\\ x & 0 < x < 1\\ 1 & x \ge 1 \end{cases}$ x0 1

Let $X \sim \text{Unif}(a, b)$. Then

$$E(X) = \frac{a+b}{2} = \text{center of interval} \qquad \xrightarrow{f_X(x)} \\ \xrightarrow{1 \\ b-a} \uparrow \\ a \\ b \\ a \\ b \\ b \\ center f_X(x) \\ ce$$

x

Let $X \sim \text{Unif}(a, b)$. Then

$$E(X) = \frac{a+b}{2} = \text{center of interval}}{2}$$

$$Var(X) = \frac{(b-a)^2}{12}$$

$$SD(X) = \frac{b-a}{2\sqrt{3}} \propto \text{ length of interval}} \xrightarrow{a \qquad b} x$$

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Proof.

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$$\operatorname{Var}(X) = E[(X - \frac{a+b}{2})^2] = \frac{1}{b-a} \int_a^b (x - \frac{a+b}{2})^2 dx = \frac{1}{b-a} \int_{(a-b)/2}^{(b-a)/2} y^2 dy = \frac{1}{b-a} \frac{2}{3} (\frac{b-a}{2})^3 = \frac{(b-a)^2}{12}$$

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 <u>Interpretation</u>: You go to lunch with your friend, say Tom, who will arrive at some time uniformly distributed between noon and 2pm. Suppose he calls you and says he will arrive by 1pm. Then his arrival time is uniform between noon and 1pm.

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- <u>Mathematically</u>: conditioned on the event $\{X \le 1\}$, X is uniformly distributed on (0, 1).

Exponential distribution

A continuous random variable X is said to be exponentially distributed with parameter $\lambda > 0$, denoted by $X \sim \text{Expo}(\lambda)$, if it has the following PDF:



Exponential distribution is often used to model time:

- The time until a radioactive particle decays
- Interarrival time of calls to a call center
- Time till the next accident, etc...

Qualitative effect of the parameter



Observation

- $\lambda \uparrow \Longrightarrow$ PDF concentrated near zero and decays faster
- $\lambda \downarrow \Longrightarrow$ PDF more spread out and decays slower

CDF

$$F_X(x) = \begin{cases} 1 - e^{-\lambda x} & x \ge 0\\ 0 & x < 0 \end{cases}$$



CDF



• From PDF to CDF:

$$F_X(x) = \int_{-\infty}^x f_X(t) dt = \int_0^x \lambda e^{-\lambda t} dt = -e^{-\lambda t} \Big|_0^x = 1 - e^{-\lambda x}$$

Tail probability:

$$P(X > x) = 1 - F_X(x) = e^{-\lambda x}$$

which becomes smaller if λ increases

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Recall integration by parts: $\int_{a}^{b} f(x) \overline{g'(x)} dx = f(x) g(x) \Big|_{a}^{b} - \int_{a}^{b} f'(x) g(x) dx.$

$$E(X) = \int_0^\infty x \lambda e^{-\lambda x} dx = \underbrace{-x e^{-\lambda x} \Big|_0^\infty}_0 + \underbrace{\int_0^\infty e^{-\lambda x} dx}_{1/\lambda} = \frac{1}{\lambda}$$

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Recall integration by parts: $\int_{a}^{b} f(x)g'(x)dx = f(x)g(x)\Big|_{a}^{b} - \int_{a}^{b} f'(x)g(x)dx.$



and

$$E(X^2) = \int_0^\infty x^2 \lambda e^{-\lambda x} dx = \underbrace{-x^2 e^{-\lambda x}}_0 \Big|_0^\infty + \underbrace{\int_0^\infty 2x e^{-\lambda x} dx}_{2E(X)/\lambda} = \frac{2}{\lambda^2}$$

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Memoryless property of geometric distribution

Recall from Lec 8: $L \sim \text{Geom}(p)$

$$P(L=k+\ell|L\geq k)=P(L=\ell),\quad k,\ell\geq 0.$$

Interpretation: Having failed k times already, the probability that one fails another ℓ times is the same as failing ℓ times from the fresh start, as if the past is "forgotten."

Memoryless property of exponential distribution

Let $X \sim \mathsf{Expo}(\lambda)$. Then

$$P(X > s + t | X > t) = P(X > s), \quad s, t > 0$$

Proof.

$$P(X > s + t | X > t) = \frac{P(X > s + t)}{P(X > t)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda t}} = e^{-\lambda s} = P(X > s).$$

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Therefore exponential distribution can be a useful model for

- Interarrival time between spam emails, visits to a website, accidents, etc
- but not a good model for
 - Lifetime of a human being, a light bulb, etc

since it does not account for the effect of "getting worn out"

How to generate random variables?



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 - Raise revenue and tax
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Question: How to generate a random variable of a specific distribution, e.g., uniform?



Random number generator: Physical methods

- 1 Hardware-based: coin, die, roulette
- Physics-based: thermal noise, cosmic background radiation, quantum...

RAND company

A Million Random Digits with 100,000 Normal Deviates Oth Edition

by The RAND Corporation (Author)



A more serious review by the famous statistician John Tukey: https://www.jstor.org/stable/166772.

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3.14
 ★★★☆☆ Spoiler alert: 8!
Reviewed in the United States on October 31, 2016
 Verified Purchase

A very engrossing book with historical importance, it keeps you guessing until the end.

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★★★☆☆ Wait for the audiobook version

By R. Rosini on October 19, 2006 Format: Paperback

While the printed version is good, I would have expected the publisher to have an audiobook version as well. A perfect companion for one's Ipod.

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Methodology

... The random digits in this book were produced by rerandomization of a basic table generated by an electronic roulette wheel.

http://www.rand.org/pubs/monograph_reports/MR1418/index2.html
https://en.wikipedia.org/wiki/List_of_random_number_generators
https://www.wsj.com/articles/

rand-million-random-digits-numbers-book-error-11600893049

Random number generator: Computerized methods

Pseudorandom number generator (PRNG): deterministic methods whose output appears random

- Simplest method: LCG
 - $\blacktriangleright X_{n+1} = (aX_n + b) \mod m$
 - ► X₀: "seed"
- Most programming languages have a rand function based on more sophisticated algorithms and a setseed function
- For most of the applications, PRNG suffices

Uniform random variable

• For a k-bit number in binary representation:

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When k is large, this is very close to Unif(0,1)

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Many PRNG produce random variables close to uniform

Given uniform random variables, how to generate random variables with other distributions?

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B-H §5.3: "Universality of the Uniform"

Quantile transformation

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- Fact: Let $U \sim \text{Unif}(0, 1)$. Then the CDF of

$$X = F^{-1}(U)$$

is given by F.

Proof.

$$P(X \le x) = P(F^{-1}(U) \le x) = P(U \le F(x)) = F(x).$$

Conversely

• Given X with CDF F,

$$U = F(X)$$

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• To summarize:

$$X \xrightarrow[F]{F^{-1}} U$$

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• To summarize:

$$X \xleftarrow{F}{F^{-1}} U$$

• Furthermore: given X with CDF F, generate Y with CDF G?



that is

$$Y = G^{-1}(F(X))$$

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Given $X \sim \text{Expo}(1)$, how to generate Expo(3):

• Recall CDF of $\text{Expo}(\lambda)$: $1 - e^{-\lambda x}$. Then

$$F(x) = 1 - e^{-x}$$

$$G(x) = 1 - e^{-3x}, \quad G^{-1}(u) = -\frac{1}{3}\ln(1 - u)$$

• Assembling everything:

$$Y = G^{-1}(F(X)) = -\frac{1}{3}\ln(1 - (1 - e^{-X})) = X/3$$

$$g(u) = \begin{cases} 0 & 0 < u < 1 - p \\ 1 & 1 - p < u < 1 \end{cases}$$



• Then $X = g(U) \sim \text{Bern}(p)$

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• Then
$$X = g(U) \sim \text{Bern}(p)$$

• Verify:

$$P(X = 0) = P(g(U) = 0) = P(0 < U < 1 - p) = 1 - p$$
$$P(X = 1) = P(g(U) = 1) = P(1 - p < U < 1) = p$$

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- This g can be viewed as inverse of CDF¹
- Clearly such g is not unique

¹For F with jumps, we define $F^{-1} = \min\{x : F(x) \ge u\}$.