## Homework 3 Due: May 12, 2016 Spring 2016 ECE 598 Information-theoretic methods in high-dimensional statistics Prof. Yihong Wu

## Rules:

- It is mandatory to type your solutions in LATEX. Email your solution in pdf (file name: HW3-your name.pdf) by midnight of the due date to yihongwu@illinois.edu with subject line Homework XX: your name.
- Justify your work rigorously. As long as you are able to prove the result or a stronger version, there is no need to follow the hints.
- 1. (Suboptimality of MLE in high dimensions) Consider the *p*-dimensional GLM:  $X \sim \mathcal{N}(\theta, I_p)$ , where  $\theta$  belongs to the parameter space

$$\Theta = \left\{ \theta \in \mathbb{R}^p : |\theta_1| \le p^{1/4}, \|\theta_{\setminus 1}\|_2 \le 2(1 - p^{-1/4}|\theta_1|) \right\}.$$

where  $\theta_{1} = (\theta_2, \ldots, \theta_p)$ . For the mean-square error loss, show that for sufficiently large p,

(a) the minimax risk is bounded:

$$\inf_{\hat{\theta}} \sup_{\theta \in \Theta} \mathbb{E}_{\theta} [\|\hat{\theta} - \theta\|_2^2] \lesssim 1$$

(b) the worst-case risk of maximal likelihood estimator

$$\theta_{\text{MLE}} \triangleq \operatorname*{arg\,min}_{\tilde{ heta} \in \Theta} \|X - \tilde{ heta}\|_2$$

is unbounded, namely,

$$\sup_{\theta \in \Theta} \mathbb{E}_{\theta} [\|\hat{\theta}_{\mathrm{MLE}} - \theta\|_2^2] \gtrsim \sqrt{p}.$$

- (c) (Bonus) Can you construct a cleaner example?
- 2. (Distribution estimation) Given n independent samples  $X_1, \ldots, X_n$  drawn from a distribution P over [k], the goal is to show that the minimax rate for estimating P with respect to the total variation loss is given by:

$$\inf_{\hat{P}} \sup_{P \in \mathcal{M}_k} \mathbb{E}_P[d_{\mathrm{TV}}(\hat{P}, P)] \asymp \sqrt{\frac{k-1}{n}} \wedge 1, \qquad k, n \in \mathbb{N},$$
(1)

where  $\mathcal{M}_k$  denotes all distributions on [k].

- (a) Show that the maximal likelihood estimator  $P_{\rm MLE}$  coincides with the empirical distribution.
- (b) Show that MLE is rate-optimal, namely, achieving the RHS of (1) within constant factor.
- (c) Establish the minimax lower bound via Assouad's lemma.
- (d) Establish the minimax lower bound via Fano's inequality + volume method or explicit packing.
- (e) (Bonus) Establish the minimax lower bound via mutual information method.

- 3. (Hard thresholding)
  - (a) Show that the solution to the  $\ell_0$ -penalized least squares

$$\hat{\theta} = \operatorname*{arg\,min}_{\theta \in \mathbb{R}^p} \|y - \theta\|_2^2 + \lambda \|\theta\|_0$$

coincides with the hard-thresholding estimator  $\hat{\theta}^{\text{HT}}$ , where

$$\hat{\theta}_i^{\mathrm{HT}} = \begin{cases} y_i & |y_i| > \tau \\ 0 & |y_i| \le \tau. \end{cases}$$

Find the relation between  $\tau$  and  $\lambda$ .

(b) Show that  $\hat{\theta}^{\text{HT}}$  is a minimizer for

$$\hat{\theta} = \operatorname*{arg\,min}_{\theta: \|y-\theta\|_{\infty} \le \tau} \|\theta\|_{0}.$$

- 4. (Soft thresholding)
  - (a) Show that the solution to the  $\ell_1$ -penalized least squares

$$\hat{\theta} = \operatorname*{arg\,min}_{\theta \in \mathbb{R}^p} \|y - \theta\|_2^2 + \lambda \|\theta\|_1$$

coincides with the soft-thresholding estimator  $\hat{\theta}^{ST}$ , where

$$\hat{\theta}_i^{\text{ST}} = \begin{cases} y_i - \tau & y_i > \tau \\ 0 & |y_i| \le \tau \\ y_i + \tau & y_i < -\tau \end{cases}$$

Find the relation between  $\tau$  and  $\lambda$ .

(b) Show that

$$\hat{\theta} = \operatorname*{arg\,min}_{\theta: \|y-\theta\|_{\infty} \le \tau} \|\theta\|_{1}$$

also coincides with  $\hat{\theta}^{\text{ST}}$ .

5. (Sharp minimax rate in sparse denoising) For *p*-dimensional GLM model  $X \sim \mathcal{N}(\theta, I_p)$ , we show that minimax risk for denoising 1-sparse vector in high dimension is

$$\inf_{\hat{\theta}} \sup_{\|\theta\|_0 \le 1} \mathbb{E}_{\theta}[\|\hat{\theta} - \theta\|_2^2] = (2 + o(1))\log p, \qquad p \to \infty.$$
(2)

(a) (Bayesian lower bound) Consider the prior  $\pi$  under which  $\theta$  is uniformly distributed over  $\{\tau e_1, \ldots, \tau e_p\}$ , where  $e_i$ 's denote the standard basis. Let  $\tau = \sqrt{(2-\epsilon)\log p}$ . Show that for any  $\epsilon > 0$ , the Bayes risk (MMSE) is given by

$$\inf_{\hat{\theta}} \mathbb{E}_{\theta \sim \pi}[\|\hat{\theta} - \theta\|_2^2] = (2 + o(1))\log p, \qquad p \to \infty.$$

(Hint: use any method you prefer, e.g., mutual information, compute the exact conditional mean and conditional variance, etc.)

(b) Demonstrate an estimator  $\hat{\theta}$  that achieves the RHS of (2) asymptotically. (Hint: one idea is to use soft thresholding with an appropriately chosen threshold).