

Fall 2018  
**Homework 2**  
**S&DS 684: Statistical Inference on Graphs**

Due: December 19, 2018

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Rules:

- It is mandatory to type your solutions in L<sup>A</sup>T<sub>E</sub>X. Email your solution in pdf by midnight of the due date to [yihong.wu@yale.edu](mailto:yihong.wu@yale.edu) with subject line **Homework XX: your name**.
  - Justify your work rigorously. As long as you are able to prove the result or a stronger version, there is no need to follow the hints.
1. (Spiked Wigner model) Consider the following rank-one perturbation to a Gaussian random matrix:

$$W = \sqrt{\frac{\mu}{n}} \sigma \sigma^\top + Z$$

where  $Z = (Z_{ij})$  is a symmetric matrix with  $\{Z_{ij} : 1 \leq i \leq j \leq n\}$  being iid  $N(0, 1)$ , and the membership vector  $\sigma$  is uniformly drawn from the set of all bisections, i.e.,  $\{\sigma \in \{\pm 1\}^n : \sum_i \sigma_i = 0\}$ .

- (a) (Detection) Consider the hypothesis testing problem of testing  $H_0 : W = Z$  (i.e.  $\mu = 0$ ) versus  $H_1 : W = \sqrt{\frac{\mu}{n}} \sigma \sigma^\top + Z$ . Assume that  $\mu$  is a constant. Show that reliable detection (i.e. both Type-I and Type-II error probabilities vanish as  $n \rightarrow \infty$ ) is impossible if  $\mu < 1$ .<sup>1</sup> (Hint: compute the  $\chi^2$ -divergence using the second moment method).
- (b) (Correlated recovery) We say an estimator  $\hat{\sigma} = \hat{\sigma}(W)$  achieves correlated recovery, if it has a nontrivial overlap with the true partition, i.e.,  $\mathbb{E}|\langle \sigma, \hat{\sigma} \rangle| = \Omega(n)$  as  $n \rightarrow \infty$ . Instead of using conditional second-moment argument as we did in class, we show that correlated recovery is impossible if  $\mu < 1$  by a reduction argument:<sup>2</sup>

Suppose correlated recovery is possible. Let's construct a test statistic. Write  $\sigma = \begin{bmatrix} \sigma_1 \\ \sigma_2 \end{bmatrix}$ , where  $\sigma_1 \in \{\pm 1\}^{(1-\epsilon)n}$  and  $\sigma_2 \in \{\pm 1\}^{\epsilon n}$  with appropriately chosen  $\epsilon$ . Write  $W$  accordingly in a block form  $W = \begin{bmatrix} W_{11} & W_{12} \\ W_{21} & W_{22} \end{bmatrix}$ . Apply correlated recovery estimator on  $W_{11}$  to obtain  $\hat{\sigma}_1$ , and compute  $y = W_{21} \hat{\sigma}_1$ . Under the null, we expect the variance of each coordinate of  $y$  is roughly 1; under the alternative, thanks to the correlation between  $\sigma_1$  and  $\hat{\sigma}_1$ , we expect the variance of each coordinate is strictly bigger than 1. Make this argument rigorous by analyzing the test statistic  $\frac{1}{n} \|y\|_2^2$ .

- (c) (Almost exact recovery) We say an estimator  $\hat{\sigma} = \hat{\sigma}(W)$  achieves almost exact recovery, if the fraction of misclassification is vanishing, i.e.,  $\mathbb{E}|\langle \sigma, \hat{\sigma} \rangle| = n - o(n)$  as  $n \rightarrow \infty$ . Show that almost exact recovery is possible if and only if  $\mu \rightarrow \infty$ . (Hint: for achievability, consider spectral method and perturbation bound).
- (d) (Exact recovery: impossibility) We say an estimator  $\hat{\sigma} = \hat{\sigma}(W)$  achieves exact recovery, if  $\mathbb{P}[\sigma = \pm \hat{\sigma}] \rightarrow 1$  as  $n \rightarrow \infty$ . Show that exact recovery is impossible if  $\mu = \sqrt{(2 - \epsilon) \log n}$  for any fixed  $\epsilon > 0$ . (Hint: show that even the maximum likelihood estimator fails in this case).

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<sup>1</sup>In fact,  $\mu = 1$  is also impossible, but we have to resort more advanced techniques than second moment method.

<sup>2</sup>This idea was suggested by Prof. Zhou Fan.

- (e) (Exact recovery: SDP) Consider the following SDP relaxation:

$$\hat{X} = \arg \max \{ \langle W, X \rangle : X \succeq 0, X_{ii} = 1, \langle X, \mathbf{J} \rangle = 0 \}$$

where  $\mathbf{J}$  is the all-one matrix. Show that exact recovery is achieved, i.e.,  $\hat{X} = \sigma\sigma^\top$  with probability tending to one, if  $\mu = \sqrt{(2 + \epsilon) \log n}$  for any fixed  $\epsilon > 0$ .

(Hint: do not invoke the general result from the lecture; instead, do a direct analysis based on two facts (i)  $\|Z\|_{op} = O(\sqrt{n})$  with high probability; (ii) the maximum of  $n$  iid standard normals is  $\sqrt{(2 + o(1)) \log n}$  with high probability).

2. ( $\|\cdot\|_{2 \rightarrow 1}$ -norm) Denote the rows of  $B \in \mathbb{R}^{n \times d}$  by  $b_1^\top, \dots, b_n^\top$ .

- (a) Show that the induced norm  $\|B\|_{2 \rightarrow 1}$  is given by

$$\|B\|_{2 \rightarrow 1} = \max \left\{ \sum_{i=1}^n |\langle b_i, y \rangle| : y \in S^{d-1} \right\}.$$

- (b) Suppose  $b_1, \dots, b_n$  are iid uniformly drawn from the sphere  $S^{d-1}$ . Show that for any fixed  $d$ , as  $n \rightarrow \infty$ ,  $\frac{1}{n} \|B\|_{2 \rightarrow 1}$  converges in probability to some value  $c_d$  as a function of  $d$ . Find  $c_d$  as explicitly as you can.

(Hint: take an  $\epsilon$ -net over  $S^{d-1}$  and use union bound. For fixed  $y \in S^{d-1}$ , what is  $\mathbb{E}[|\langle b_1, y \rangle|]$ ?)

- (c) Show that  $\sqrt{d}c_d \rightarrow \sqrt{\frac{2}{\pi}}$  as  $d \rightarrow \infty$ .

3. (Grothendieck inequality for PSD matrices) For  $A = (a_{ij}) \in \mathbb{R}^{n \times n}$ , consider

$$\|A\|_{\infty \rightarrow 1} \triangleq \max \left\{ \sum_{i,j \in [n]} a_{ij} x_i y_j : x_i, y_j \in \{\pm 1\} \right\} = \max \left\{ \langle A, xy^\top \rangle : \|x\|_\infty \leq 1, \|y\|_\infty \leq 1 \right\} \quad (1)$$

and its SDP relaxation

$$\text{SDP}(A) \triangleq \max \left\{ \sum_{i,j \in [n]} a_{ij} \langle u_i, v_j \rangle : u_i, v_j \in S^{n-1} \right\} = \max \{ \langle A, X \rangle : X \succeq 0, X_{ii} = 1 \}. \quad (2)$$

- (a) Following the argument in class, show that for every positive semidefinite  $A$ ,

$$\text{SDP}(A) \geq \|A\|_{\infty \rightarrow 1} \geq \frac{2}{\pi} \text{SDP}(A). \quad (3)$$

- (b) Next we show that the constant  $\frac{2}{\pi}$  in (3) is sharp by constructing instances of  $A$  so that the ratio  $\frac{\|A\|_{\infty \rightarrow 1}}{\text{SDP}(A)}$  is arbitrarily close to  $\frac{2}{\pi}$ .

- (i) Show that without loss of optimality, we can restrict to  $x_i = y_i$  in (1);  
(Hint:  $\langle A, xy^\top \rangle^2 = \langle \sqrt{A}x, \sqrt{A}y \rangle^2 \leq \langle A, xx^\top \rangle \langle A, yy^\top \rangle$ . Why?)
- (ii) Show that without loss of optimality, we can restrict to  $u_i = v_i$  in (2);  
(Hint:  $\langle A, U^\top V \rangle^2 = \langle \sqrt{A}U^\top, \sqrt{A}V^\top \rangle^2 \leq \langle A, U^\top U \rangle \langle A, V^\top V \rangle$ . Why?)

(iii) Show the following deterministic fact: if  $A = \frac{1}{n^2}BB^\top$  for  $B \in \mathbb{R}^{n \times d}$ , then

$$\|A\|_{\infty \rightarrow 1} = \frac{1}{n^2} \|B\|_{2 \rightarrow 1}^2$$

and

$$\text{SDP}(A) \geq \frac{1}{n^2} \sum_{i,j \in [n]} \langle b_i, b_j \rangle^2$$

where  $b_i^\top$  is the  $i$ th row of  $B$ .

- (iv) Now take the rows of  $B$  to be iid uniform on  $S^{d-1}$ , use the previous problem to show that  $\|A\|_{\infty \rightarrow 1} \rightarrow c_d^2$  as  $n \rightarrow \infty$ .
- (v) Show that  $\text{SDP}(A) \geq r_d$  in probability as  $n \rightarrow \infty$ , where  $r_d$  behaves as  $\frac{1+o(1)}{d}$  as  $d \rightarrow \infty$ .
- (vi) Conclude the sharpness of the constant  $\frac{2}{\pi}$  in (3).  
(Hint: what is  $\mathbb{E}[\langle b_1, b_2 \rangle^2]$ ? Be careful with  $\sum_{i,j \in [n]} \langle v_i, v_j \rangle^2$  which is not an iid sum.)